

Longitudinal trimmed and smooth trimmed effects with flip and S-flip interventions

Alec McClean and Iván Díaz

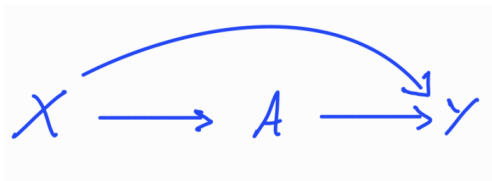
New York University Grossman School of Medicine



Trimming with single timepoint data

$\{(X_i, A_i, Y_i)\}_{i=1}^n \stackrel{iid}{\sim} \mathcal{P}$ with X_i covariates, $A_i \in \{0, 1\}$ binary treatment and $Y_i \in \mathbb{R}$ outcome.

$Y(a)$ is potential outcome under treatment a .



Positivity violations: $\pi(X) = \mathbb{P}(A = 1 \mid X) \approx 0$ or ≈ 1 .

$\mathbb{E}\{Y(a)\}$ not identified ($\pi = 0$) or hard-to-estimate ($\pi \approx 0$)

Trimming: Change population to satisfy positivity

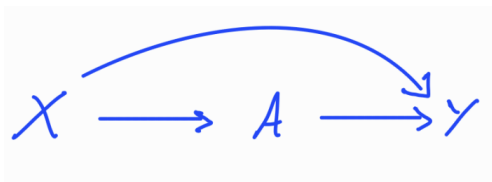
$$\mathbb{E}\left[\{Y(1) - Y(0)\} \mathbb{1}\{\varepsilon < \pi(X) < 1 - \varepsilon\}\right] \text{ for } \varepsilon \in [0, 0.5)$$

(related to cond. trimming: $\mathbb{E}\{f(Z) \mid \mathcal{E}\} = \mathbb{E}\{f(Z)\mathbb{1}(\mathcal{E})\}/\mathbb{P}(\mathcal{E})$)

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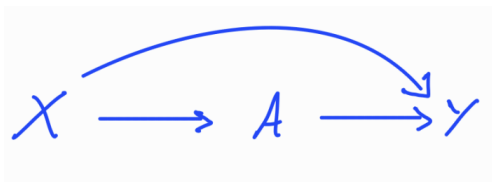
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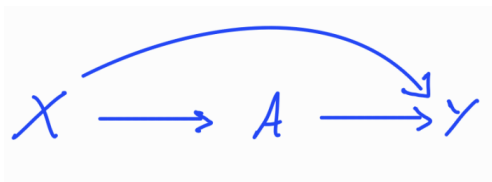
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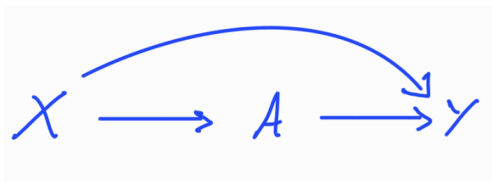
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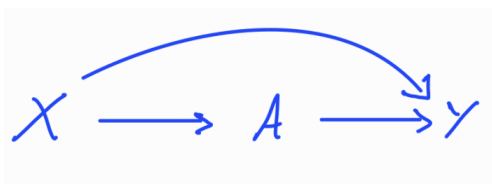
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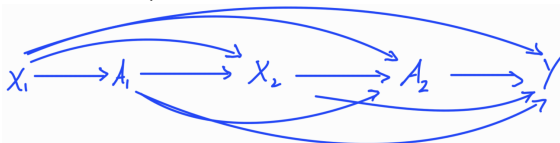
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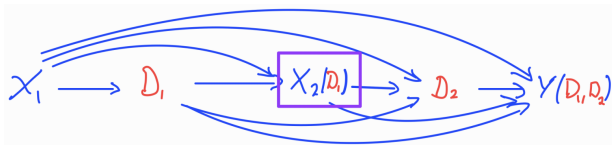
Trimming on non-baseline covariates is challenging

Data: (X_1, A_1, X_2, A_2, Y) .



Can we trim on propensity scores in the second timepoint?
(where we often have positivity violations!)

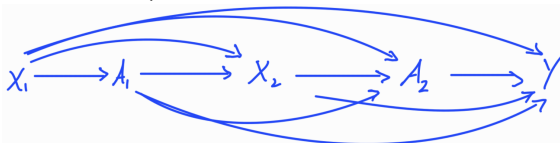
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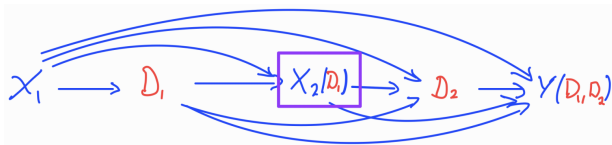
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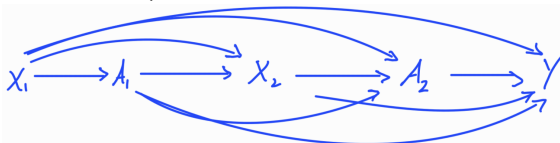
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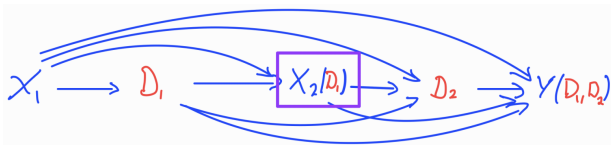
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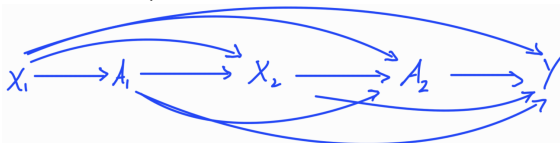
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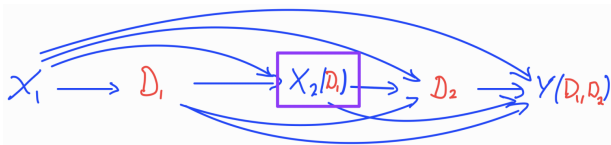
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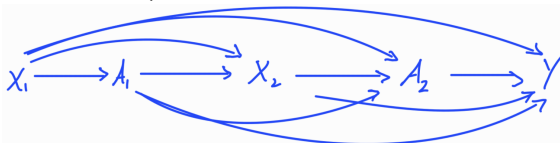
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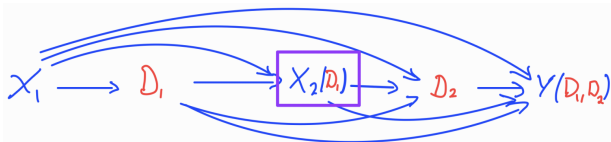
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Trimming is a dynamic “flip” intervention

Dynamic stochastic interventions:

$D = d(A, X)$ adapts so $\mathbb{E}\{Y(D)\}$ IDable and estimable

- ▶ Modified treatment policies [Díaz et al., 2023]
- ▶ Incremental propensity score interventions [Kennedy, 2019]

Proposition

Let $\mathbb{1}(X) \equiv \mathbb{1}\{\varepsilon < \pi(X) < 1 - \varepsilon\}$ and

$$D(a) = \begin{cases} A & \text{if } A = a \\ \mathbb{1}(X)a + \{1 - \mathbb{1}(X)\}A & \text{if } A \neq a. \end{cases}$$

Then, if $\{Y(1), Y(0)\} \perp\!\!\!\perp A \mid X$,

$$\mathbb{E}[Y\{D(1)\} - Y\{D(0)\}] = \mathbb{E}[\{Y(1) - Y(0)\}\mathbb{1}\{\varepsilon < \pi(X) < 1 - \varepsilon\}]$$

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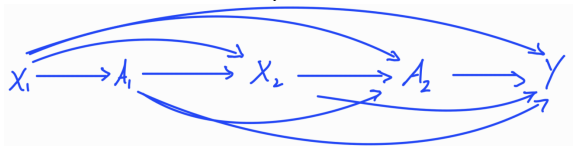
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Longitudinal setup

Longitudinal data

$\{Z_i\}_{i=1}^n \stackrel{iid}{\sim} \mathbb{P} \in \mathcal{P}$ where $Z = (X_1, A_1, X_2, A_2, \dots, X_T, A_T, Y)$



$X_t \in \mathbb{R}^d$: time-varying covariates

$A_t \in \{0, 1\}$: time-varying binary treatment

$Y \in \mathbb{R}$: ultimate outcome of interest

History of O_t at t : $\overline{O}_t = (O_1, \dots, O_t)$

Future of O_t from t : $\underline{O}_t = (O_t, \dots, O_T)$

$H_t = (\overline{X}_t, \overline{A}_{t-1})$: covariate and treatment history at t

NPSEM assumption: there are $\{f_{X,t}, f_{A,t}\}_{t=1}^T$ and f_Y such that

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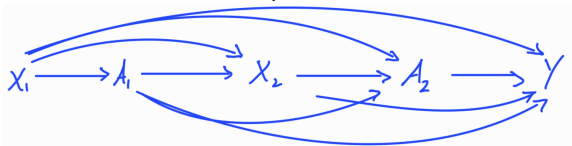
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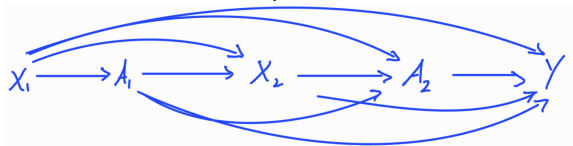
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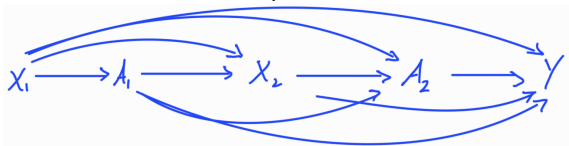
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Causal assumptions

NPSEM embeds **consistency**

We will **avoid positivity** by **trimming/flipping**
(e.g. $\exists \delta > 0$ s.t. $\mathbb{P}\{\delta < \mathbb{P}(A_t = 1 \mid H_t) < 1 - \delta\} = 1.$)

Sequential randomization:

- 1 **Standard:** $U_{A,t} \perp\!\!\!\perp (\underline{U}_{X,t+1}, U_Y) \mid H_t$ for all $t \leq T$
- 2 **Strong:** $U_{A,t} \perp\!\!\!\perp (\underline{U}_{X,t+1}, \underline{U}_{A,t+1}, U_Y) \mid H_t$

Standard SR: common causes of A_t and future covariates and outcome measured. Typical in literature.

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- ② **Strong:** $U_{A,t} \perp\!\!\!\perp (\underline{U}_{X,t+1}, \underline{U}_{A,t+1}, U_Y) \mid H_t$

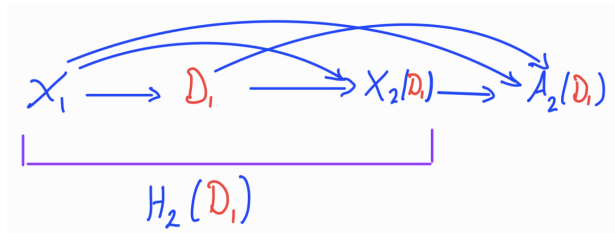
Standard SR: common causes of A_t and future covariates and outcome measured. Typical in literature.

Strong SR: common causes A_t and future covariates and outcome **and treatments** measured. Allows for ID when intervention depends on **natural value of treatment**.

Counterfactual variables under $\bar{D}_{t-1}$

$\bar{D}_{t-1}$ generates **counterfactual variables at time t** :

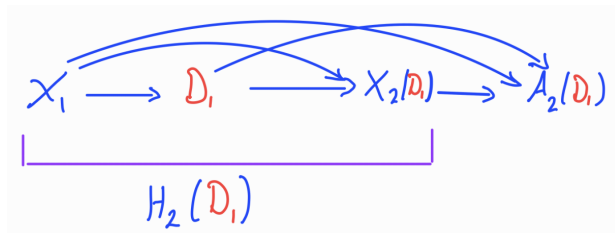
- ▶ $X_t(\bar{D}_{t-1}) = f_{X,t}(D_{t-1}, H_{t-1}(\bar{D}_{t-2}), U_{X,t})$
"Natural covariate value"
- ▶ $H_t(\bar{D}_{t-1}) = (\bar{D}_{t-1}, \bar{X}_t(\bar{D}_{t-1}))$
"Natural covariate and treatment history"
- ▶ $A_t(\bar{D}_{t-1}) = f_{A,t}(H_t(\bar{D}_{t-1}), U_{A,t})$
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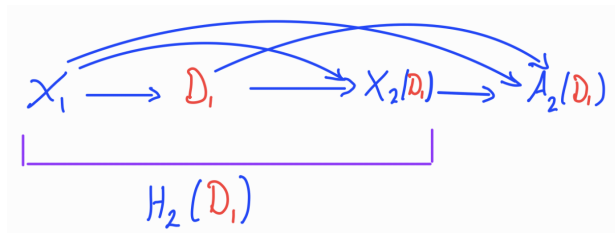
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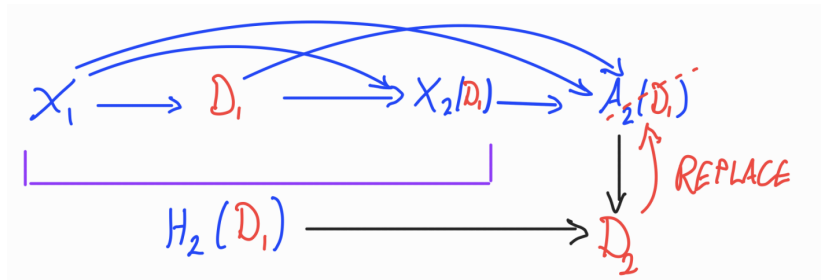
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Intervention D_t and potential outcomes $\bar{D}_T$

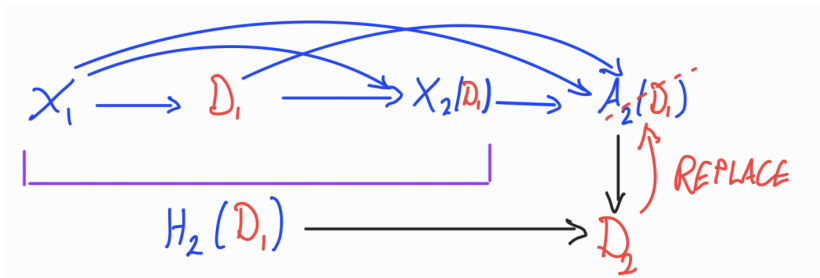


D_t can be a function of $A_t(\bar{D}_{t-1})$, $H_t(\bar{D}_{t-1})$.

Ultimately, replace $\bar{A}_T$ with $\bar{D}_T$:

- $Y(\bar{D}_T) = f_Y(D_T, H_T(\bar{D}_{T-1}), U_Y)$
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Longitudinal trimming with flip interventions

Longitudinal flip interventions

- ▶ $\bar{a}_T \in \{0, 1\}^T$ is the target regime
- ▶ $I_t^c \equiv I_t^c(a_t; h_t) = \mathbb{1}[\mathbb{P}\{A_t(\bar{D}_{t-1}) = a_t \mid H_t(\bar{D}_{t-1}) = h_t\} > \varepsilon]$ is trimming indicator

Flip intervention at time t targeting a_t is

$$D_t(a_t) = \begin{cases} A_t(\bar{D}_{t-1}), & \text{if } A_t(\bar{D}_{t-1}) = a_t, \\ I_t^c a_t + (1 - I_t^c) A_t(\bar{D}_{t-1}), & \text{otherwise,} \end{cases}$$

At t , if the natural value of treatment is **already** a_t , **do nothing**; otherwise, “**flip**” the subject to a_t if their **natural history lies in the trimmed set** (determined by I_t^c).

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Theorem: identification

Let $I_t(a_t; h_t) = \mathbb{1}\{\mathbb{P}(A_t = a_t \mid h_t) > \varepsilon\}$. Suppose

- ▶ $\bar{D}_T = \{D_1(a_1), D_2(a_2), \dots, D_T(a_T)\}$ are **flip interventions** and
- ▶ the **NPSEM** and **strong SR** hold.

Then,

$$\begin{aligned} \mathbb{E}\{Y(\bar{D}_T)\} &= \sum_{\bar{b}_T \in \{0,1\}^T} \mathbb{E} \left\{ \overbrace{\mathbb{E}(Y \mid \bar{A}_T = \bar{b}_T, \bar{X}_T) \prod_{t=1}^T Q_t(b_t \mid \bar{b}_{t-1}, \bar{X}_t)}^{\text{extended g-formula}} \right\}, \\ &= \mathbb{E} \left[Y \prod_{t=1}^T \frac{Q_t(A_t \mid H_t)}{\mathbb{P}(A_t \mid H_t)} \right] \text{IPW}, \end{aligned}$$

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$$Q_t(a_t \mid h_t) = I_t(a_t; h_t) + \{1 - I_t(a_t; h_t)\} \mathbb{P}(A_t = a_t \mid h_t)$$

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We are trimming, but we can't(/shouldn't) aim for “trimmed” effects

Not necessarily true that $\mathbb{E}\{Y(\overline{D}_T) - Y(\overline{D}'_T)\}$ isolates

$$\psi = \mathbb{E}\left[\{Y(\overline{a}_T) - Y(\overline{a}'_T)\} \mathbb{1}(\text{could take both } \overline{a}_T \text{ and } \overline{a}'_T)\right].$$

Why? because $\overline{D}_t$ affects $X_{t+1}(\overline{D}_t), A_{t+1}(\overline{D}_t)$

However, ψ is a bad target!

Theorem: Other flip interventions $\overline{D}_T$ and $\overline{D}'_T$ yield $\mathbb{E}\{Y(\overline{D}_T) - Y(\overline{D}'_T)\} = \psi$. However, they have two properties:

- 1 They depend on cross-world propensity scores: trim supposing non-zero propensity score under $D_1 = 1$ and $D_1 = 0$
non-falsifiable
- 2 They depend on future propensity scores: trim at $t = 1$ supposing non-zero propensity scores at $t = 2$

(this justifies the caution typically advised)

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Smooth trimming with S-flip
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Smooth trimming

Everything we've done so far can be generalized to smooth trimming!

Why: flip effects inherit $\hat{\pi}$ convergence rate

ML estimator $\implies$ slower-than- $\sqrt{n}$ convergence

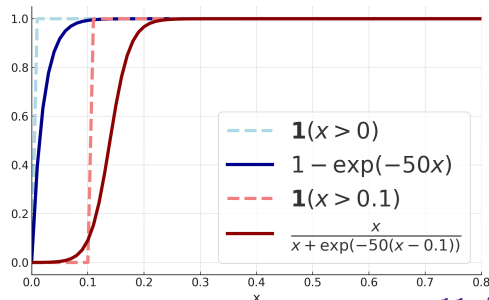
Smooth trimmed effects: $\sqrt{n}$ -convergence w/ ML estimators

How: replace $\mathbb{1}(\cdot)$ by $S(\cdot)$: $\mathbb{E}[\{Y(1) - Y(0)\}S(X)]$

Examples:

① $\mathbb{1}(x > 0) \rightarrow$
 $s(x; k) = 1 - \exp(-kx)$

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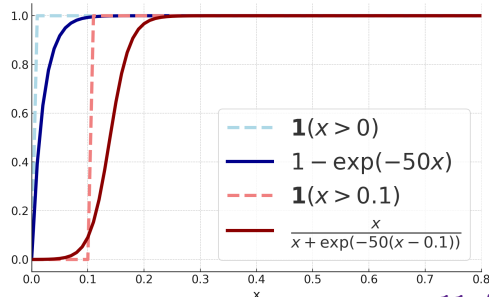
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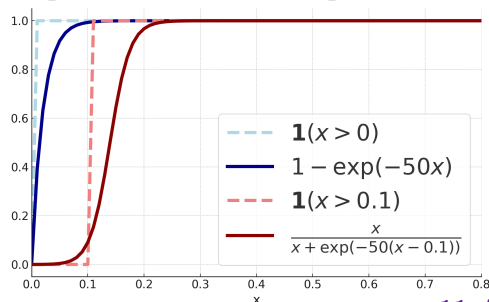
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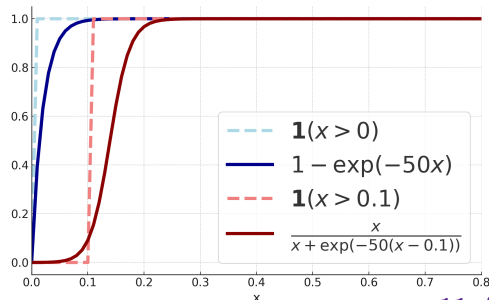
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and $V_1, \dots, V_T \stackrel{iid}{\sim} \text{Uniform}(0, 1)$ and $V_t \perp\!\!\!\perp Z$ for all $t \leq T$.

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$s(x) = 1 - \exp(-50x)$ approximates $\mathbb{1}(x > 0)$:
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Smooth-flip (S-flip) interventions

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Theorem: S-flip effect identification

Let $S_t(a_t; h_t) = s\{\mathbb{P}(A_t = a_t \mid h_t), \varepsilon, k\}$. Suppose

- ▶ $\bar{D}_T = \{D_1(a_1), D_2(a_2), \dots, D_T(a_T)\}$ are **S-flip interventions**,
- ▶ $s(0) = 0$, and
- ▶ the **NPSEM** and **strong SR** hold.

Then,

$$\begin{aligned}\mathbb{E}\{Y(\bar{D}_T)\} &= \sum_{\bar{b}_T \in \{0,1\}^T} \mathbb{E}\left\{\mathbb{E}(Y \mid \bar{A}_T = \bar{b}_T, \bar{X}_T) \prod_{t=1}^T Q_t(b_t \mid \bar{b}_{t-1}, \bar{X}_t)\right\}, \\ &= \mathbb{E}\left[Y \prod_{t=1}^T \frac{Q_t(A_t \mid H_t)}{\mathbb{P}(A_t \mid H_t)}\right],\end{aligned}$$

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Estimation

(in two slides)

Estimation of S-flip effects

S-flip effects are pathwise differentiable $\implies$ **efficient influence function** (EIF)-based estimators

We derive the new EIF (plug-in plus weighted residuals)

Inspires two one-step estimators:

- 1 Multiply robust-style
- 2 Sequentially doubly robust-style (debias pseudo-outcome in multiply robust-style estimator)

We derive bias bounds, which imply weak convergence for S-flip estimator under nonparametric conditions on nuisance functions

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Bias bounds

Two new results for multiply robust-style estimator:

- 1 Observe that bias is **minimum** of two errors
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→ Hence, $2(T + 1)$ multiply robust-style bound
- 2 Tighten bias bound from IPSI-style result [Kennedy, 2019]

Sequentially DR-style result is **first-of-its-kind** where Q_t depends on **unknown propensity score**

Theorem, informal: Let ψ denote S-flip effect, π_t prop. score at t , $\tilde{m}_t$ denote seq. reg. at t with estimated pseudo-outcome. Then,

$$\left| \mathbb{E} \left(\hat{\psi}_{sdr} - \psi \right) \right| \lesssim \sum_{t=1}^T \|\hat{\pi}_t - \pi_t\| \left(\|\hat{m}_t - \tilde{m}_t\| + \|\hat{\pi}_t - \pi_t\| \right).$$

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$\overbrace{\|\hat{\pi}_t - \pi_t\| = o_{\mathbb{P}}(n^{-1/4}) = \|\hat{m}_t - \tilde{m}_t\|}^{\text{ML conv. rates}} \implies \sqrt{n}\text{-consistency \& asymp-}$
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Thank you!

Recap:

- ▶ $T = 1$:
trimming $\equiv$ flip interventions
smooth trimming $\equiv$ S-flip ints
- ▶ $T > 1$: not so simple.
Flip/S-flip ints avoid positivity
asmp and cross-world /
future-dependence
- ▶ Efficient estimation of S-flip
effects: (1) multiply robust and
(2) sequentially doubly robust

Ongoing work:

- ▶ Data analysis (administrative
censoring and censoring by
death)

hadera01@nyu.edu

Prelim draft:



[alecmcclean.github.io/
files/LSTTEs-short.pdf](https://alecmcclean.github.io/files/LSTTEs-short.pdf)

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- Aksel KG Jensen, Theis Lange, Olav L Schjørring, and Maya L Petersen. Identification and responses to positivity violations in longitudinal studies: an illustration based on invasively mechanically ventilated icu patients. *Biostatistics & Epidemiology*, 8(1):e2347709, 2024.
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- Alexander W Levis, Edward H Kennedy, Alec McClean, Sivaraman Balakrishnan, and Larry Wasserman. Stochastic interventions, sensitivity analysis, and optimal transport. *arXiv preprint arXiv:2411.14285*, 2024.

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Additional remarks

- ① Robustness to positivity violations $\implies$ **S-flips are alternative to IPSIs that target a specific regime**
- ② **Relax dependence on natural value of treatment:**

$$D_t(a_t) = \mathbb{1} \left[V_t \leq \mathbb{1}(a_t = 1)I_t^c + (1 - I_t^c) \mathbb{P}\{A_t(\overline{D}_{t-1}) = 1 \mid H_t(\overline{D}_{t-1})\} \right]$$

- ▶ Can **avoid practical issues** — we may not observe natural value of treatment
- ▶ Identification **only requires standard sequential randomization**
- ③ Flip interventions inspired by **maximally coupled policies** Levis et al. [2024]